

# Evaluation of Nowcasting for Detecting and Predicting Local Influenza Epidemics, Sweden, 2009–2014

Armin Spreco, Olle Eriksson, Örjan Dahlström, Benjamin John Cowling, Toomas Timpka

The growing availability of big data in healthcare and public health opens possibilities for infectious disease control in local settings. We prospectively evaluated a method for integrated local detection and prediction (nowcasting) of influenza epidemics over 5 years, using the total population in Östergötland County, Sweden. We used routine health information system data on influenza-diagnosis cases and syndromic telenursing data for July 2009–June 2014 to evaluate epidemic detection, peak-timing prediction, and peak-intensity prediction. Detection performance was satisfactory throughout the period, except for the 2011–12 influenza A(H3N2) season, which followed a season with influenza B and pandemic influenza A(H1N1) pdm09 virus activity. Peak-timing prediction performance was satisfactory for the 4 influenza seasons but not the pandemic. Peak-intensity levels were correctly categorized for the pandemic and 2 of 4 influenza seasons. We recommend using versions of this method modified with regard to local use context for further evaluations using standard methods.

Although the seasonal variations in influenza incidence among nations and global regions are well described (1), the duration and intensity of influenza epidemics in local communities have been less well monitored and understood. The rapidly growing availability of big data from diagnostic and prediagnostic (syndromic) sources in healthcare and public health settings opens new possibilities for increasing the granularity in infectious disease control (2,3). However, development of outbreak models and efficient use of the information produced by prediction models in public health response decision-making remain challenging. This observation was recently highlighted

by the Congress of the United States request that the Government Accountability Office gather information on validation of emerging infectious disease model predictions (<https://energycommerce.house.gov/wp-content/uploads/2017/11/20171109GAO.pdf>).

We previously reported the design of a nowcasting method (detection of influenza epidemics and short-term predictions) for local-level application in the northwestern region of the world (4). In other fields, such as meteorology, nowcasting methods represent standard tools for warning the public against dangerous high-impact events (5). The rationale for developing this novel method was to inform the planning of local responses and adjustments of healthcare capacities. Many such adjustments are planned and performed locally, at county and municipality levels. In Sweden, for instance, the hospital bed capacity is habitually overextended; on average, 103 patients occupy 100 regular hospital bed units (6). It is therefore important that an influenza epidemic is noticed early at the local level to make time for implementation of adjustments (e.g., freeing hospital beds by removing from the waiting list those patients scheduled for elective interventions).

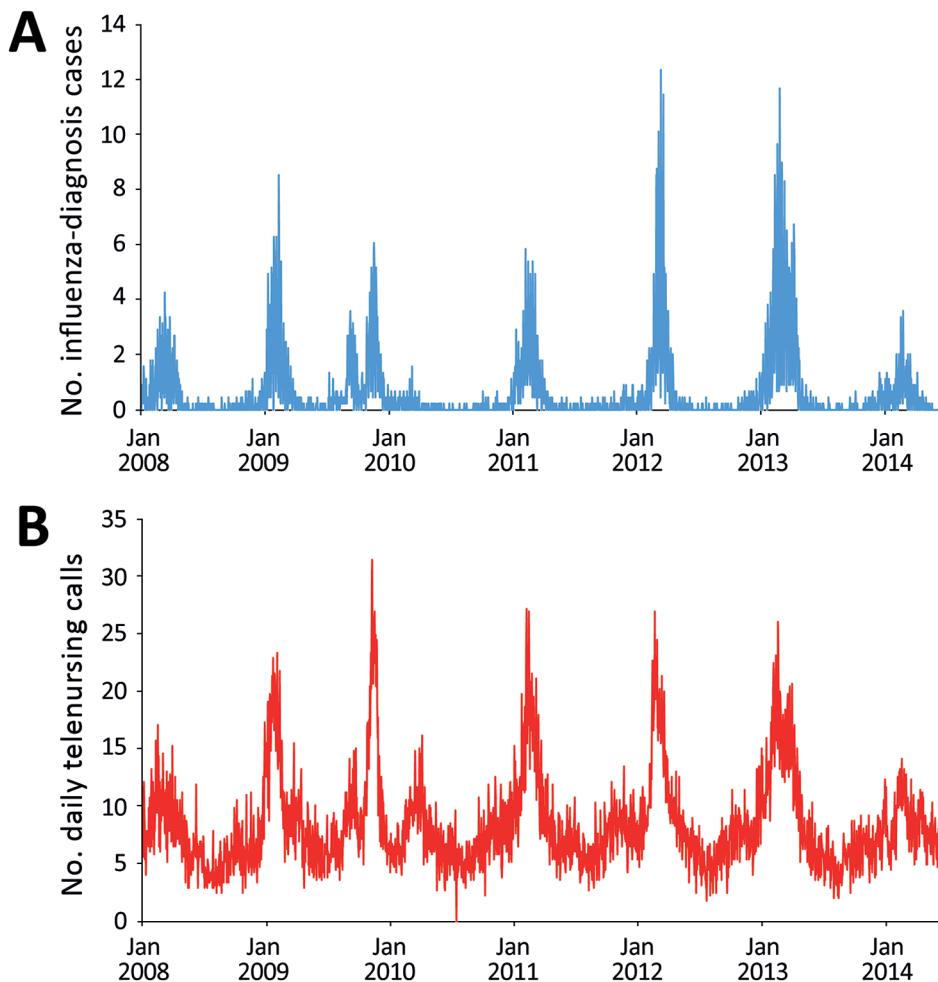
We performed a prospective 5-year evaluation of local influenza nowcasting by using routine health information system data. The evaluation period included 1 pandemic (2009) and 4 winter influenza seasons (Figure). The nowcasting method is based on mathematical modeling of epidemic curves generated from historic local data (4). Nowcasting comprises 3 functions: detection of the local start of the epidemic, prediction of peak timing, and prediction of peak intensity.

## Methods

We used an open cohort design based on the total population ( $n = 445,000$ ) in Östergötland County, Sweden. We used prospective data from July 1, 2009, through June 30, 2014, from 2 sources in the countywide health information system: clinical influenza-diagnosis cases recorded by physicians and syndromic chief complaint data from a

Author affiliations: Linköping University, Linköping, Sweden (A. Spreco, O. Eriksson, Ö. Dahlström, T. Timpka); Center for Health Services Development, Region Östergötland, Linköping (A. Spreco, T. Timpka); Hong Kong University, Hong Kong Special Administrative Region, China (B.J. Cowling)

DOI: <https://doi.org/10.3201/eid2410.171940>



**Figure.** Data used for evaluation of nowcasting for detection and prediction of local influenza epidemics, Östergötland County, Sweden, January 1, 2008, through June 30, 2014.

A) Unadjusted daily numbers of influenza-diagnosis cases per 100,000 population.

B) Unadjusted daily telenursing calls because of fever (child and adult) per 100,000 population.

telenursing service (4). The influenza-diagnosis case data were used for detection of the local start of the epidemic and prediction of its peak intensity; the syndromic data were used for prediction of the peak timing. Timeliness was used as a performance metric for detection of the local start of the epidemic and the peak-timing prediction; the correct identification of intensity category on a 5-grade scale was used for peak-intensity prediction. The study design was approved by the Regional Research Ethics Board in Linköping (no. 2012/104–31).

### Definitions

We identified influenza cases by using codes from the International Classification of Diseases, 10th Revision, for influenza (J10.0, J10.1, J10.8, J11.0, J11.1, J11.8) (7) as recorded in the local electronic health data repository. We identified influenza-related telenursing call cases by using the chief complaint codes associated with influenza symptoms (dyspnea, fever [child and adult], cough [child and adult], sore throat, lethargy, syncope, dizziness, and headache [child and adult]).

The intensity level for the start of a local epidemic (i.e., the endpoint for the detection function) was set to 6.3 influenza-diagnosis cases/100,000 population recorded during a floating 7-day period in the countywide health information system (3). A recent comparison of influenza intensity levels in Europe estimated a similar definition (6.4 influenza-diagnosis cases/wk/100,000 population) for the 2008–09 winter influenza season in Sweden (8). Peak timing was defined as the date when the highest number of influenza-diagnosis cases were documented in the countywide electronic patient record. Peak intensity was defined as the number of influenza-diagnosis cases that had been documented on that date.

The detection threshold was adjusted to situations when extended simmering of influenza-related activity appears before an epidemic. Such simmering was associated with exceptionally mild winter influenza seasons and pandemics. Preepidemic simmering before winter influenza seasons was defined as occurring when the period between increased influenza incidence above baseline and the start of the epidemic is prolonged (4). The upper

threshold for the baseline was set to 3.2 influenza-diagnosis cases/100,000 population during a floating 7-day period (i.e., half of the defined start-of-epidemic intensity level). If the baseline threshold was surpassed for a period 3 times longer than the average period before previous epidemics without exceeding the start-of-epidemic level (i.e., 6.3 influenza-diagnosis cases/100,000 population during a floating 7-day period), preepidemic simmering was considered to have occurred. Preepidemic simmering in association with a pandemic was defined as that ensuing from the date of a World Health Organization pandemic alert. We determined all definitions and adjustments before using the method for detection and prediction.

### Method Application

Technical details concerning the 3 functions of nowcasting are provided in online Technical Appendix 1 (<https://wwwnc.cdc.gov/EID/article/24/10/17-1940-Techapp1.pdf>). The programming code for the analyses is provided in online Technical Appendix 2 (<https://wwwnc.cdc.gov/EID/article/24/10/17-1940-Techapp2.pdf>).

To calibrate the detection component of the nowcasting method, we retrospectively determined weekday effects on recording of influenza-diagnosis cases and a baseline alarm threshold by using learning data. These data were collected from January 1, 2008, through June 30, 2009, including the 2 winter influenza seasons 2007–08 and 2008–09. To determine weekday effects, we used data from the entire learning data collection period. To determine the initial alerting threshold, we used only data from the winter influenza season in 2008–09. The 2007–08 winter influenza season could not be used for this purpose because collection of learning data had already started. Throughout the study period, the calibration data were updated after every winter influenza season (i.e., no updates after the 2009 pandemic outbreak). The detection algorithm was thus applied to the next epidemic by using the revised threshold determined in the updated learning dataset. We identified 2 exceptional situations: pandemic settings and winter influenza seasons that simmered before they started (4). In these situations, the alarm threshold is doubled. Accordingly, following the World Health Organization pandemic alert (9), the alarm threshold was doubled for the 2009 season. Before the 2010–11 winter influenza season, the threshold was reset to the regular level. No updates were performed because the set of learning data remained the same (i.e., it contained data from the 2008–09 winter influenza season). For the 2011–12 winter influenza season, we updated the threshold by using learning data from the 2008–09 and 2010–11 winter influenza seasons. For the 2012–13 winter influenza season, we updated the threshold by using learning data from the 2008–09, 2010–11, and 2011–12 winter influenza seasons. For the 2013–14 winter influenza season, we again

updated the alerting threshold by using learning data from the previous winter influenza seasons (2008–09, 2010–11, 2011–12, and 2012–13). However, because this winter influenza season simmered before it started, the threshold was doubled before the detection method was applied. The weekday effects were assumed to be relatively constant over time in the local detection analyses and therefore were not updated after every winter influenza season.

We also used the set of learning data from the winter influenza seasons in 2007–08 and 2008–09 to initially calibrate the first component of the local prediction module. The dataset was used to decide the grouping of chief complaints with the largest correlation strength and longest lead time between influenza-diagnosis data and telenursing data (10,11). The best performing telenursing chief complaint was fever (child and adult), and the most favorable lead time was 14 days. When the peak timing had been determined, the second component of the local prediction module was applied to influenza-diagnosis data from the corresponding epidemics to find the peak intensity on the predicted peak day (4). Regarding weekday effects on local prediction, the same calculation was applied and the same grouping of chief complaints and lead time were used throughout the study.

### Metrics and Interpretations

For trustworthiness of the nowcasting method in local healthcare planning, we set the maximum acceptable timeliness error for detection and peak timing predictions to 1.5 weeks. Method performance was defined to be excellent if the absolute value of the timeliness error was  $\leq 3$  days, good if it was 4–7 days, acceptable if it was 8–11 days, and poor if it was  $\geq 12$  days. For peak intensity predictions, we used the epidemic threshold and intensity level categories (nonepidemic, low, medium, high, and very high) identified for Sweden in a study involving 28 European countries (8) (Table 1). If the predicted peak intensity fell into the same category as the actual peak intensity, the prediction was considered successful; otherwise, it was considered unsuccessful.

## Results

### Local Detection

The detection component of the local nowcasting method showed good performance during the 2009 pandemic of influenza A(H1N1)pdm09 (pH1N1) virus (Table 2), alerting for the local influenza epidemic 5 days after it actually started. For the 2010–11 winter influenza season with influenza B and pH1N1 viruses circulating, the local detection performance was also good; the alarm was raised 5 days after the start of the local epidemic. During the 2011–12 winter influenza season, with influenza A(H3N2) virus

**Table 1.** Epidemic intensity categories used to interpret performance measurements in evaluation of nowcasting for detection and prediction of local influenza epidemics, Sweden, 2009–2014\*

Intensity level	Threshold, cases/d/100,000 population					
	2008–09	2009 pandemic	2010–11	2011–12	2012–13	2013–14
Nonepidemic	<0.9	<0.9	<0.9	<1.0	<1.2	<1.2
Low	0.9	0.9	0.9	1.0	1.2	1.2
Medium	2.4	2.5	2.5	2.5	2.8	2.9
High	5.5	5.4	5.4	5.2	5.6	5.5
Very high	7.9	7.5	7.5	7.1	7.7	7.4

\*Based on (8).

activity only, detection performance was poor. During that season, the alarm was raised 15 days before the actual local epidemic started. During the 2012–13 and 2013–14 winter influenza seasons, with influenza A(H3N2), influenza B, and pH1N1 viruses circulating, the local detection performance was excellent; alarms were raised 3 days before the 2012–13 epidemic started and 3 days after the 2013–14 epidemic started.

### Local Prediction

For the 2009 influenza pandemic, performance of the local peak-timing prediction was poor, but the peak-intensity level was correctly categorized (medium intensity epidemic) (Table 3). For the 2010–11 winter influenza season, with influenza B and pH1N1 viruses circulating, the local peak-timing prediction was excellent and the peak intensity was successfully predicted to the correct category (medium intensity). For the 2011–12 winter influenza seasons with influenza A(H3N2) virus circulating and the 2012–13 winter season with influenza A(H3N2), influenza B, and pH1N1 viruses circulating, the local peak-timing predictions were good. The local peak-intensity predictions were successful for the 2012–13 winter influenza season, correctly categorizing it to a very high intensity level and unsuccessful for the 2011–12 season, categorizing it as a medium intensity epidemic when it actually developed into a very high-intensity epidemic. For the 2013–14 winter influenza season, with influenza B and pH1N1 viruses circulating and a simmering start, the local peak-timing prediction was acceptable, but the local peak intensity was wrongly predicted to be at the nonepidemic level when the winter influenza season actually reached a medium intensity level.

### Discussion

In this prospective 5-year evaluation of a method for local nowcasting of influenza epidemics that used routine health information system data, we identified aspects that were satisfactory and identified areas where improvements are needed. The detection function displayed satisfactory performance throughout the evaluation period, except for the 2011–12 winter influenza season, in which influenza A(H3N2) virus circulated after a season with influenza B and pH1N1 virus activity. Peak-timing prediction performance was satisfactory for the 4 winter influenza seasons but not for the 2009 pandemic. In addition, the method categorized the local peak-intensity levels correctly for the 2009 pandemic and for 2 of the winter influenza seasons, but it was unsuccessful at forecasting the very high peak intensity of the 2011–12 season and the medium peak intensity of the 2013–14 season, which was preceded by a simmering phase.

The results indicate that securing the availability of a new data source is only the first step toward using the data stream in routine surveillance. The syndromic data source used for the big data stream in this study was subjected to rigorous restructuring and maintenance. Nonetheless, not all parameters associated with the syndromic data stream were regularly updated. For the peak-timing predictions made by using telenursing data, we assumed that increases in telenursing activity precede influenza diagnoses by 14 days. Although this assumption is grounded (10,11), the interval may change over time and thereby influence influenza predictions. Using the constant interval estimate, we estimated the influenza diagnosis peaks for the 2011–12 and 2012–13 winter influenza seasons 1 week before and 1 week after the actual influenza-diagnosis peaks. In

**Table 2.** Performance of the detection algorithm displayed with alert thresholds updated by using data from previous winter influenza seasons in evaluation of nowcasting for detection and prediction of local influenza epidemics, Sweden, 2009–2014\*

Influenza virus activity	Updated threshold,		Interpretation
	cases/d/100,000 population	Timeliness†	
2009 pH1N1‡	0.424	–5	Good
2010–11 B and pH1N1	0.212	–5	Good
2011–12 A(H3N2)	0.207	15	Poor
2012–13 A(H3N2), B, and pH1N1	0.242	3	Excellent
2013–14 A(H3N2), B, and pH1N1‡	0.481	–3	Excellent

\*pH1N1, pandemic influenza A(H1N1)pdm09 virus.

†Positive value means that the algorithm issued an alarm before the local epidemic had started; negative value means that the alarm was raised after the start of the epidemic.

‡The threshold was doubled because of a pandemic alert or observation of a period of simmering influenza activity.

**Table 3.** Performance of peak-timing and peak-intensity predictions from evaluation of nowcasting for detection and prediction of local influenza epidemics, Sweden, 2009–2014\*

Influenza virus active	Time-of-peak predictions†				Peak intensity predictions		
	Date when prediction made	Time to peak, d	Prediction error	Interpretation	Category (cases/d/100,000 population)		Interpretation
					Predicted	Factual	
2009 pH1N1	2009 Sep 27	8	–28	Poor	Medium (3.3)	Medium (2.9)	Successful
2010–11 B and pH1N1	2011 Feb 11	10	0	Excellent	Medium (4.5)	Medium (4.9)	Successful
2011–12 A(H3N2)	2012 Feb 25	9	7	Good	Medium (4.5)	Very high (12.4)	Unsuccessful
2012–13 A(H3N2), B, and pH1N1	2013 Feb 22	10	–7	Good	Very high (10.1)	Very high (11.7)	Successful
2013–14 A(H3N2), B, and pH1N1	2014 Feb 17	8	–8	Acceptable	Nonepidemic (1.0)	Medium (3.4)	Unsuccessful

\*pH1N1, pandemic influenza A(H1N1)pdm09 virus.

†Positive value means that the peak was predicted to be reached before the actual peak occurred; negative value means that the peak was predicted to be reached after the actual peak occurred.

other words, the basic prediction method may have been more accurate at predicting the peak timing than the results of this study show. In the setting of our study, the performance was most likely decreased by the assumption that telenursing precedes influenza diagnosis by 14 days is applicable to all situations. Althouse et al. reported that methods underpinning the use of big data sources (e.g., search query logs) need regular upkeep to maintain their accuracy (12). In future versions of nowcasting methods, regular updates and syndromic sources that are more stable than telenursing data (regarding time lag to influenza diagnosis data) may become available and can be used to improve the peak-timing predictions.

Influenza forecasting is methodologically challenging (13,14), and only a few prospective evaluations have transparently reported algorithms and study designs. In the first Centers for Disease Control and Prevention (CDC) challenge, a prospective study of state-of-the-art methods in which 4 aspects of influenza epidemics (start week, peak week, peak percentage, and duration) were forecasted by using routine data (15), none of the evaluated methods showed satisfactory performance for all aspects. Similarly, in the second CDC challenge, in which 3 aspects of influenza epidemics (start week, peak week, and peak intensity) were forecasted by using 7 methods, none of the evaluated methods displayed satisfactory performance (16). These challenge studies have substantially helped to widen the understanding of the difficulties of forecasting different aspects of influenza epidemics. In our study, the detection and peak-intensity prediction functions of the local nowcasting method underperformed during the 2011–12 winter influenza season. One reason for the observed underperformance may be that the data from preceding seasons used to generate local epidemic curves were insufficient for modeling the between-seasons drift in the immunity status of the population in relation to the circulating influenza strain. In other words, the present parameters used to compute

epidemic curves were deficient when large drifts in population immunity with corresponding changes in virus dissemination patterns occurred. For instance, the epidemic phase of the influenza A(H3N2) season in 2011–12 may not have started with virus spread among the young persons in the community as it had during previous seasons (17,18). It has been suggested that including virologic information (i.e., influenza virus type and subtype) as model parameters may improve the predictive accuracy of mathematical models (19). We contend also that, for local influenza detection and prediction, historical accounts of the circulating influenza virus types should be considered for inclusion in the statistical models and suggest adding information about the population age structure. However, such model extensions must also be paralleled by securing a continuous supply of the corresponding data in the local settings where the models are to be used.

This study has strengths and weaknesses that need to be considered when interpreting the results. The main strength of the study is that it prospectively evaluates an integrated influenza nowcasting method in a local community. On the basis of experiences from previous studies (4,15,16,20), we considered timeliness to be the most valid general evaluation metric for our purposes. To be able to accurately support adjustments of local healthcare capacity, we used daily data for the evaluations. In the CDC challenge studies (15,16), forecasts of the start and peak timing of an epidemic were based on weekly data and considered accurate if they occurred within 1 week of the actual timing of each component. Therefore, we consider the limits used for evaluating detection and the peak-timing predictions in this study to be at least as strict as those in the CDC challenge. Regarding the prediction of peak intensity, we considered a forecast to be accurate if it predicted the peak intensity to be the correct peak-intensity category as defined by Vega et al. (8), who calculated the thresholds for each of these categories for every winter influenza season by applying

the moving epidemic method (21) on 5–10 previously occurring seasons. Hence, we consider these categories to be reliable. A longer prospective evaluation period would have increased the possibility of drawing valid conclusions concerning the outcome of the evaluation, and it would be preferable to have corresponding local data from other cities or regions (22). Evaluating our nowcasting method for epidemics from several other regions would enable conclusions to be drawn about the generalizability of the method.

We contend that methods for local nowcasting of influenza epidemics based on routine health information system data have potential for general dissemination and use. Future versions of the nowcasting model will be gradually extended with information on population age distribution and on current and previously circulating influenza types. Such extensions need to be paralleled by securing a routine supply of data to the added parameters in local health information systems. We recommend using versions of the nowcasting method modified with regard to their local use context for further evaluations with standard measures.

This study was supported by grants from the Swedish Civil Contingencies Agency (2010–2788) and the Swedish Research Council (2008–5252). The funders had no role in the study design, data collection and analysis, decision to publish, or preparation of the manuscript.

### About the Author

Dr. Spreco is a researcher in the field of syndromic surveillance at Linköping University, Sweden. His main research focus is on evaluation and development of algorithms for detection and prediction of infectious diseases.

### References

- Polansky LS, Outin-Blenman S, Moen AC. Improved global capacity for influenza surveillance. *Emerg Infect Dis*. 2016;22:993–1001. <http://dx.doi.org/10.3201/eid2206.151521>
- Shaman J, Karspeck A, Yang W, Tamerius J, Lipsitch M. Real-time influenza forecasts during the 2012–2013 season. *Nat Commun*. 2013;4:2837. <http://dx.doi.org/10.1038/ncomms3837>
- Riley RD, Ensor J, Snell KI, Debray TP, Altman DG, Moons KG, et al. External validation of clinical prediction models using big datasets from e-health records or IPD meta-analysis: opportunities and challenges. *BMJ*. 2016;353:i3140. <http://dx.doi.org/10.1136/bmj.i3140>
- Spreco A, Eriksson O, Dahlström Ö, Cowling BJ, Timpka T. Integrated detection and prediction of influenza activity for real-time surveillance: algorithm design. *J Med Internet Res*. 2017;19:e211. <http://dx.doi.org/10.2196/jmir.7101>
- Blišňák V, Sokol Z, Zacharov P. Nowcasting of deep convective clouds and heavy precipitation: comparison study between NWP model simulation and extrapolation. *Atmos Res*. 2017;184:24–34. <http://dx.doi.org/10.1016/j.atmosres.2016.10.003>
- Sveriges Kommuner och Landsting. Ingen på sjukhus i onödan. No one in hospital unnecessarily [in Swedish]. Stockholm: Sveriges Kommuner och Landsting; 2016.
- World Health Organization. International statistical classification of diseases and related health problems. 10th Revision, vol. 2. Geneva: The Organization; 2010.
- Vega T, Lozano JE, Meerhoff T, Snacken R, Beauté J, Jorgensen P, et al. Influenza surveillance in Europe: comparing intensity levels calculated using the moving epidemic method. *Influenza Other Respi Viruses*. 2015;9:234–46. <http://dx.doi.org/10.1111/irv.12330>
- World Health Organization. Influenza A(H1N1) [cited 2017 Jan 15]. [http://www.who.int/mediacentre/news/statements/2009/h1n1\\_20090429/en/](http://www.who.int/mediacentre/news/statements/2009/h1n1_20090429/en/)
- Timpka T, Spreco A, Dahlström Ö, Eriksson O, Gursky E, Ekberg J, et al. Performance of eHealth data sources in local influenza surveillance: a 5-year open cohort study. *J Med Internet Res*. 2014;16:e116. <http://dx.doi.org/10.2196/jmir.3099>
- Timpka T, Spreco A, Eriksson O, Dahlström Ö, Gursky EA, Strömngren M, et al. Predictive performance of telenursing complaints in influenza surveillance: a prospective cohort study in Sweden. *Euro Surveill*. 2014;19:46. <http://dx.doi.org/10.2807/1560-7917.ES2014.19.46.20966>
- Althouse BM, Scarpino SV, Meyers LA, Ayers JW, Bargsten M, Baumbach J, et al. Enhancing disease surveillance with novel data streams: challenges and opportunities. *EPJ Data Sci*. 2015;4:pii:17. <http://dx.doi.org/10.1140/epjds/s13688-015-0054-0>
- Chretien JP, George D, Shaman J, Chitale RA, McKenzie FE. Influenza forecasting in human populations: a scoping review. *PLoS One*. 2014;9:e94130. <http://dx.doi.org/10.1371/journal.pone.0094130>
- Lazer D, Kennedy R, King G, Vespignani A. Big data. The parable of Google Flu: traps in big data analysis. *Science*. 2014;343:1203–5. <http://dx.doi.org/10.1126/science.1248506>
- Biggerstaff M, Alper D, Dredze M, Fox S, Fung IC, Hickmann KS, et al.; Influenza Forecasting Contest Working Group. Results from the Centers for Disease Control and Prevention's predict the 2013–2014 influenza season challenge. *BMC Infect Dis*. 2016;16:357. <http://dx.doi.org/10.1186/s12879-016-1669-x>
- Biggerstaff M, Johansson M, Alper D, Brooks LC, Chakraborty P, Farrow DC, et al. Results from the second year of a collaborative effort to forecast influenza seasons in the United States. *Epidemics*. 2018 Feb 24;pii:S1755–4365(17)30088–9. Epub ahead of print [cited 2018 Jun 14]. <http://dx.doi.org/10.1016/j.epidem.2018.02.003>
- Schanzer D, Vachon J, Pelletier L. Age-specific differences in influenza A epidemic curves: do children drive the spread of influenza epidemics? *Am J Epidemiol*. 2011;174:109–17. <http://dx.doi.org/10.1093/aje/kwr037>
- Timpka T, Eriksson O, Spreco A, Gursky EA, Strömngren M, Holm E, et al. Age as a determinant for dissemination of seasonal and pandemic influenza: an open cohort study of influenza outbreaks in Östergötland County, Sweden. *PLoS One*. 2012;7:e31746. <http://dx.doi.org/10.1371/journal.pone.0031746>
- Moniz L, Buczak AL, Baugher B, Guven E, Chretien JP. Predicting influenza with dynamical methods. *BMC Med Inform Decis Mak*. 2016;16:134. <http://dx.doi.org/10.1186/s12911-016-0371-7>
- Cowling BJ, Wong IO, Ho LM, Riley S, Leung GM. Methods for monitoring influenza surveillance data. *Int J Epidemiol*. 2006;35:1314–21. <http://dx.doi.org/10.1093/ije/dyl1162>
- Vega T, Lozano JE, Meerhoff T, Snacken R, Mott J, Ortiz de Lejarazu R, et al. Influenza surveillance in Europe: establishing epidemic thresholds by the moving epidemic method. *Influenza Other Respir Viruses*. 2013;7:546–58. <http://dx.doi.org/10.1111/j.1750-2659.2012.00422.x>
- Moss R, Zarebski A, Dawson P, McCaw JM. Retrospective forecasting of the 2010–2014 Melbourne influenza seasons using multiple surveillance systems. *Epidemiol Infect*. 2017;145:156–69. <http://dx.doi.org/10.1017/S0950268816002053>

Address for correspondence: Armin Spreco, Linköping University, Division of Social Medicine, Department of Medical and Health Sciences, Faculty of Health Sciences, SE-581 83 Linköping, Sweden; email: armin.spreco@liu.se

# Evaluation of Nowcasting for Detecting and Predicting Local Influenza Epidemics, Sweden, 2009–2014

## Technical Appendix 1

### Method Design Overview

The nowcasting method is divided into separate modules for detection and prediction of influenza activity, respectively (*I*). An overview of the main statistical assumptions and equations for each method component is displayed in the Technical Appendix Figure. The function of the detection module is to alert for an upcoming period of increased load of influenza-diagnosis cases on local health care services, whereas the function of the prediction module is to predict the timing of the activity peak and its intensity. The prediction process is divided into two components. In the first component, syndromic data are used to predict the peak timing, and in the second component, influenza-diagnosis data are used to estimate the peak intensity.

In the study setting, patients clinically diagnosed with influenza were used as gold standard. Early detection of increased influenza activity and prediction of peak intensity were thus based on streams of the gold standard data, whereas prediction of peak timing was based on syndromic data.

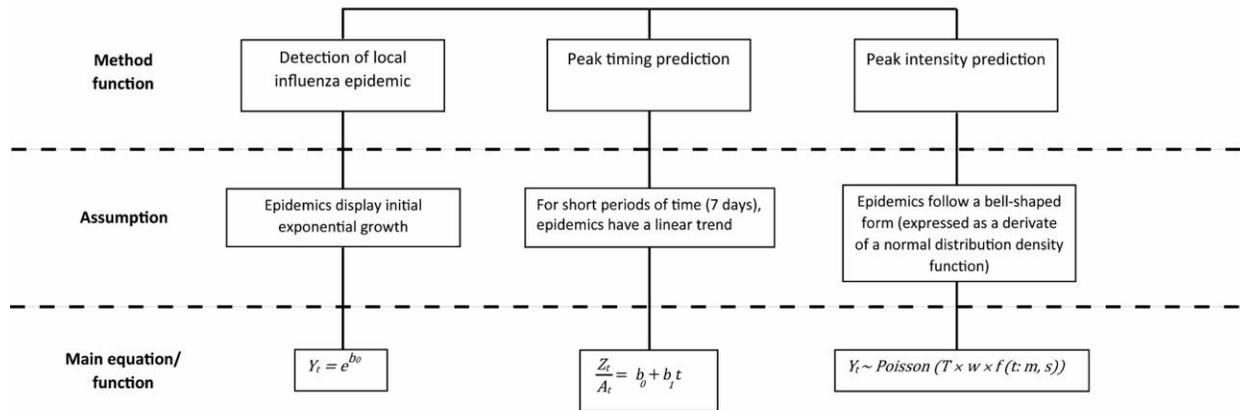


Figure. Overview of the main mathematical equations or functions used for each component.

## Detection Module

Exponential regression (1) is used for detection modeling, based on the observation that the beginning of an influenza epidemic is assumed to have an exponential growth of infected individuals;

$$(1) X_t = e^{a_0 + b_1 t},$$

with  $t$  representing the time,  $a_0$  representing the level, and  $b_1$  representing the trend. The expected number of visits at local health care services,  $E[Y_t]$  is the product of  $X$  and the probability  $p$  for an infected individual to visit the local health care service. This expectation is also exponential in time;

$$(2) E[Y_t] = e^{a_0 + b_1 t} p = e^{a_0 + \ln(p) + b_1 t} = e^{b_0 + b_1 t},$$

Where  $b_0$  now combines the current level of number of infected and probability of visiting the local health care service without any possibility to separate them. As daily data are used in the analysis, weekday effects,  $A_w$ , are also calculated and used as an offset variable in the exponential regression analysis. The weekday effects are calculated as follows: let  $A_{Monday}$  be the average number of events on Mondays during previous epidemics and denote the values for other weekdays by  $A_{Tuesday}$ ,  $A_{Wednesday}$ , and so on. Let  $A_{Total} = (A_{Monday} + \dots + A_{Sunday})/7$ . The multiplicative

weekday effect for Mondays is  $A_{Monday}/A_{Total}$  and so on. The weekday effects are included in the model;

$$(3) E[Y_t] = e^{b_0 + b_1 t + \ln(A_w)},$$

If  $X$  is large,  $p$  is small, and the infected individuals act independently, then  $Y$  is approximately Poisson distributed;

$$(4) Y_t \sim \text{Poisson}(e^{b_0 + b_1 t + \ln(A_w)}).$$

Furthermore, the time is shifted, that is, the most recent day is considered as  $t = 0$ , the second most recent day is considered as  $t = -1$ , and so on. For every new day, the time axis is moved one step so that the new “most recent day” is considered as  $t = 0$ . For each day an exponential regression analysis (1) is run and a fitted value  $\hat{y}$  is calculated by inserting  $t = 0$  in equation (3) giving

$$(5) Y_t = e^{b_0},$$

as an estimate of the current level of visits which is smoothed for random variation and adjusted for weekday effects. This is repeated for each day by moving the time axis 1 day at a time so that the most recent point in time of the series is considered  $t = 0$ . Doing this, one value is obtained for every day representing the level for that day. Finally, the lower 95% confidence limit is calculated to represent the level of influenza activity, which is then compared with a predetermined threshold. If the level (i.e., the lower 95% confidence limit) is above the threshold, an alarm is raised, which means that the winter influenza season (or pandemic) has started; and if the level is below the threshold, no alarm is raised.

Detection starts when the previous epidemic has ended (the interepidemic period level for the community where the detection component is applied), and runs during the inter-epidemic period until an increase in diagnosed influenza cases is detected. When the increase is confirmed, the algorithm is paused and restarted when the epidemic has ended.

## Prediction Module

### Peak Timing Prediction

In the first component, the aim is to predict the peak timing using linear regression. Including weekday effects  $A_w$  and smoothed for random variation, the model for the number of cases in syndromic data are expressed as

$$(6) Z_t = (b_0 + b_1 t) \times A_w,$$

with  $b_0$  representing the level and  $b_1$  representing the trend. Since the weekday effects  $A_w$  are known, a model smoothed for weekday effects and random variation can be expressed as

$$(7) Z_t / A_w = b_0 + b_1 t$$

For each 7-day period, a linear regression (7) is run and parameter estimates  $b_0$  and  $b_1$  are fitted. The idea is to estimate the trend in syndromic data for every 7-day period (the first period being days 1–7 and the second being days 2–8), from the beginning of an epidemic and until the peak is found. Although it is unlikely that an epidemic curve increases and decreases linearly, the assumption can be made that the trend during a short period of 7 days has almost a linear increase or decrease.

The search for the peak starts when the detection algorithm signals that an epidemic has taken off and continues until the peak is detected. To identify the peak timing, two conditions are set. As per the first condition, it is essential to ensure that the epidemic has a sufficiently sharp upward trend. The trend is therefore defined as sufficiently sharp when significantly positive ( $p < 0.05$ ) trends  $b_1$  have occurred either during two consecutive or during three different 7-day periods. When one of these events has occurred, the second condition is applied. According to this condition, when the first significantly negative trend ( $b_1$ ) during a 7-day period has occurred, it is assumed that the peak has been reached on the first day of this period. However, there is a possibility that this 7-day period “overlaps” with a previous 7-day period, which includes a significantly positive trend. In that case, the first 7-day period with a significantly negative trend is ignored and the peak is instead assumed to appear during the second 7-day period with a significantly negative trend. The search is aborted if the peak is not found when the epidemic has already descended in the local setting where the algorithm is applied.

When the peak is found in the syndromic data, the 14 days preceding influenza-diagnosis data (2) is used to find the peak (in influenza-diagnosis data). In other words, if the peak in the syndromic data appears on day 0, the influenza-diagnosis peak is assumed to appear on day 14. However, it is possible that the peak in the syndromic data occurs on a day during the weekend but highly unlikely that the peak in influenza-diagnosis data occurs on one of these days as, for instance, primary care centers are closed during weekends in Sweden. Instead, it is reasonable to assume that the influenza-diagnosis peak occurs at the beginning of the week because individuals who suffer influenza symptoms during the weekend visit primary care centers when they reopen on Monday or possibly Tuesday. Adjustments are therefore made by moving the influenza-diagnosis peak to the following Monday if it is expected to occur on a Friday, Saturday, or Sunday according to syndromic data and to the previous Tuesday if the peak is expected to take place on a Wednesday or Thursday. If the peak is expected to occur on a Monday or Tuesday, no adjustments are made. In other words, in the first case the syndromic data precedes influenza-diagnosis data between 15 and 17 days, in the second case between 12 and 13 days, and in the third case 14 days.

Depending on what day of the week the peak in the syndromic data are expected to take place, the prediction of the influenza-diagnosis peak is made between 6 and 11 days before it is expected to occur, as the syndromic peak can be determined first after 6 days has passed of the syndromic data series.

### **Peak Intensity Prediction**

In the second component of the prediction module, the aim is to predict only the peak intensity. Based on empirical assessments of previous epidemics, an epidemic adjusted for weekday effects is assumed to show a bell-shaped form from the beginning to the end, and can therefore be expressed using a derivate of a normal distribution density function. The intensity function must also include weekday effects and total number of events during the whole epidemic. Assuming that the peak timing is known (estimated in the first prediction component) and that an epidemic follows the bell-shaped function around the peak, the intensity function can be used to predict the peak intensity at time  $m$ .

Assume that day number  $t = 1, 2, 3, \dots, t_i$ ; the observed number of influenza-diagnosis cases is  $y = y_1, y_2, y_3, \dots, y_i$ , and that

$$(8) Y_t \sim \text{Poisson}(T \times w \times f(t; m, s)),$$

where  $T$  is the total number of health care visits of the whole epidemic,  $w$  is the weekday effects,  $f$  is the normal distribution density function,  $t$  is the day number,  $m$  is the center of the epidemic (which coincides with  $t$  for the peak), and  $s$  is the spread in time. Since  $t$ ,  $w$ , and  $m$  are known, only the parameters  $T$  and  $s$  are estimated using  $y$  in such way so that the likelihood is maximized. However, to do that, first appropriate starting values for these parameters need to be selected. Finally, using the known parameter  $m$  and the estimated parameters  $T$  and  $s$ , the peak intensity at time  $m$  is calculated by replacing  $t$  with  $m$  in equation (8).

It is important that the start of the series seems appropriate because the second prediction component assumes that the level is zero or at an interepidemic level at the start and it is not optimal that there are single or occasional spikes at the beginning of the series. For that reason, the start of the series should be a couple of weeks before an epidemic is detected.

## References

1. Spreco A, Eriksson O, Dahlström Ö, Cowling BJ, Timpka T. Integrated detection and prediction of influenza activity for real-time surveillance: Algorithm design. *J Med Internet Res.* 2017;19:e211. [PubMed http://dx.doi.org/10.2196/jmir.7101](http://dx.doi.org/10.2196/jmir.7101)
2. Timpka T, Spreco A, Dahlström Ö, Eriksson O, Gursky E, Ekberg J, et al. Performance of eHealth data sources in local influenza surveillance: a 5-year open cohort study. *J Med Internet Res.* 2014;16:e116. [PubMed http://dx.doi.org/10.2196/jmir.3099](http://dx.doi.org/10.2196/jmir.3099)

# Evaluation of Nowcasting for Detection and Prediction of Local Influenza Epidemics, Sweden, 2009–2014

## Technical Appendix 2

### Programming code

#### Detection Module

The analyses and coding of this module were performed in IBM SPSS Statistics. The code was written using the SPSS Syntax. Descriptive text/comment of the code is added by beginning a line using an asterisk (i.e., \*) and is terminated by a period.

\* Program:

\* 1. Paste the data/time series into IBM SPSS Statistics, including the variables influenza-like-illness cases (ILI), time and logarithmic values of the weekday effects.

\* 2. The dependent variable is the ILI cases.

\* 3. The independent variable is the time, where time is “shifted,” i.e., the last point in time is equal to 0, the second last point in time is  $-1$  and so on.

\* 4. The logarithmic values of the weekday effects are used as an offset variable and these are calculated using the initial learning data.

\* 5. Run the initial program below.

GENLIN ILI WITH Time

/MODEL Time INTERCEPT = YES OFFSET = Weekdayeffect

DISTRIBUTION = POISSON LINK = LOG

/CRITERIA METHOD = FISHER (1) SCALE = 1 COVB = MODEL

MAXITERATIONS = 100 MAXSTEPHALVING = 5 PCONVERGE = 1E-006(ABSOLUTE)

SINGULAR = 1E-012 ANALYSISTYPE = 3(WALD) CILEVEL = 95 CITYPE = WALD  
LIKELIHOOD = FULL

/MISSING CLASSMISSING = EXCLUDE

/PRINT SOLUTION.

\* 6. Run the program below t-1 times, where t is the length of the time series (in days).

DEFINE !SimReg (NumLoops = !CMDEND)

!DO !i = 1 !TO !NumLoops

USE ALL.

COMPUTE filter\_\$ = (Time<0).

VARIABLE LABELS filter\_\$ 'Time<0 (FILTER)'.  
VALUE LABELS filter\_\$ 0 'Not Selected' 1 'Selected'.  
FORMATS filter\_\$ (f1.0).  
FILTER BY filter\_\$.  
EXECUTE.  
COMPUTE Time = Time+1.  
EXECUTE.  
\*Generalized Linear Models.  
GENLIN ILI WITH Time  
/MODEL Time INTERCEPT = YES OFFSET = Weekdayeffect  
DISTRIBUTION = POISSON LINK = LOG  
/CRITERIA METHOD = FISHER (1) SCALE = 1 COVB = MODEL  
MAXITERATIONS = 100 MAXSTEPHALVING = 5 PCONVERGE = 1E-006(ABSOLUTE)  
SINGULAR = 1E-012 ANALYSISTYPE = 3(WALD) CILEVEL = 95 CITYPE = WALD  
LIKELIHOOD = FULL  
/MISSING CLASSMISSING = EXCLUDE

/PRINT SOLUTION.

!DOEND

!ENDDEFINE.

\* 7. Change the number of loops below, from X to t-1.

!SimReg NumLoops = X.

\* 8. Save the lower 95% confidence limits of the intercept from each regression analysis.

\* These represents the level of the influenza activity and are compared with the predetermined threshold.

\* If the level is above the threshold, an alarm is raised, which means that the winter influenza season (or pandemic) has started; and if the level is below the threshold, no alarm is raised.

### **Prediction Module – Peak Timing Prediction**

The analyses and coding of this module were performed in IBM SPSS Statistics. The code was written using the SPSS Syntax.

\* Program:

\* 1. Paste the data/time series into IBM SPSS Statistics, including the variables telenursing cases, time and the weekday effects.

\* 2. The dependent variable is the telenursing cases adjusted for weekday effects.

\* 3. The independent variable is the time, ranging from 1 to x, depending on the length of the series.

\* 4. First the telenursing cases are adjusted for weekday effects.

COMPUTE Y\_A = Telenursing/Weekdayseffect.

EXECUTE.

\* 5. Run the initial program below for the first 7-day period.

USE ALL.

COMPUTE filter\_\$ = (Time>0&Time<8).

VARIABLE LABELS filter\_\$ 'Time>0&Time<8 (FILTER)'.  
VALUE LABELS filter\_\$ 0 'Not Selected' 1 'Selected'.

FORMATS filter\_\$ (f1.0).

FORMATS filter\_\$ (f1.0).

FILTER BY filter\_\$.

EXECUTE.

\* Generalized Linear Models.

GENLIN Y\_A WITH Time

/MODEL Time INTERCEPT = YES

DISTRIBUTION = NORMAL LINK = IDENTITY

/CRITERIA SCALE = MLE COVB = MODEL PCONVERGE = 1E-006(ABSOLUTE)

SINGULAR = 1E-012 ANALYSISTYPE = 3(WALD) CILEVEL = 95 CITYPE = WALD

LIKELIHOOD = FULL

/MISSING CLASSMISSING = EXCLUDE

/PRINT SOLUTION.

\* 6. Run the second program below for the second 7 day period, where first the 7-day period is moved one step by calculating the new variable  $Time2 = Time-1$ .

COMPUTE Time2 = Time-1.

EXECUTE.

USE ALL.

COMPUTE filter\_\$ = (Time2>0&Time2<8).

VARIABLE LABELS filter\_\$ 'Time2>0&Time2<8 (FILTER)'.  
VALUE LABELS filter\_\$ 0 'Not Selected' 1 'Selected'.

FORMATS filter\_\$ (f1.0).

FORMATS filter\_\$ (f1.0).

FILTER BY filter\_\$.

EXECUTE.

\* Generalized Linear Models.

GENLIN Y\_A WITH Time2

/MODEL Time2 INTERCEPT = YES

DISTRIBUTION = NORMAL LINK = IDENTITY

/CRITERIA SCALE = MLE COVB = MODEL PCONVERGE = 1E-006(ABSOLUTE)  
SINGULAR = 1E-012 ANALYSISTYPE = 3(WALD) CILEVEL = 95 CITYPE = WALD  
LIKELIHOOD = FULL

/MISSING CLASSMISSING = EXCLUDE

/PRINT SOLUTION.

\* 7. Run the program below t-9 times, where t is the length of the time series (in days) of interest and the number 9 covers the 2 programs already run above and the 7 days covered in the regression analysis of the (seven-day) period.

DEFINE !SimReg (NumLoops = !CMDEND)

!DO !i = 1 !TO !NumLoops

COMPUTE Time2 = Time2-1.

EXECUTE.

USE ALL.

COMPUTE filter\_\$ = (Time2>0&Time2<8).

VARIABLE LABELS filter\_\$ 'Time2>0&Time2<8 (FILTER)'.  
VALUE LABELS filter\_\$ 0 'Not Selected' 1 'Selected'.

FORMATS filter\_\$ (f1.0).

FILTER BY filter\_\$.

EXECUTE.

EXECUTE.

\*Generalized Linear Models.

```

GENLIN Y_A WITH Time2

/MODEL Time2 INTERCEPT = YES

DISTRIBUTION = NORMAL LINK = IDENTITY

/CRITERIA SCALE = MLE COVB = MODEL PCONVERGE = 1E-006(ABSOLUTE)
SINGULAR = 1E-012 ANALYSISTYPE = 3(WALD) CILEVEL = 95 CITYPE = WALD
LIKELIHOOD = FULL

/MISSING CLASSMISSING = EXCLUDE

/PRINT SOLUTION.

!DOEND

!ENDDEFINE.

```

\* 8. Change the number of loops below, from X to t-9.

```
!SimReg NumLoops = X.
```

\* 9. Save the regression coefficient for the variable time as well as the associated p-value from each regression analysis.

\* Find the timing of the peak by following the description of this module provided in Online Technical Appendix 1 (<https://wwwnc.cdc.gov/EID/article/24/10/17-1940-Techapp1.pdf>).

## Prediction Module – Peak Intensity Prediction

The analyses this module were performed in Microsoft Excel.

Program:

1. Paste the following variables in a worksheet in Microsoft Excel:

- The date in column A.
- The time  $t$  (day number) in column B, where  $t$  goes from 1 to  $m$  (representing the day of the peak timing). The starting day of the series, i.e., day 1, should be a couple of weeks before the epidemic begins.

- The number of known ILI-cases  $y_j$  per day in column C, where  $j$  goes from 1 to the last known observation, i.e., being the day when the peak timing prediction is made (between  $m-6$  and  $m-11$ , see last paragraph of “Peak Timing Prediction” in Online Technical Appendix 1).

- The number of the weekday (1 = Monday, 2 = Tuesday, ...7 = Sunday) in column D
- The weekday effects  $w$  in column E.

2. Paste/create the following parameters in the same worksheet:

- The timing of the peak  $m$  in cell K1 (assumed to be known, estimated in the peak timing prediction module).

- The total number of ILI-cases of the whole epidemic  $T$  in cell K2.  $T$  is unknown, but an appropriate starting value must be selected, for instance 100.

- The spread in time  $s$  in cell K3.  $s$  is also unknown, but an appropriate starting value must be selected also here, for instance 8.

- In cell K4, calculate the sum of the values in column J, representing the sum of the logarithmic likelihood (see below how this sum is calculated).

3. Create the following variables in the same worksheet:

- The probability density function of a Normal distribution ( $t:m,s$ ) in column F for days 1 to  $m$ , by applying the Excel function  $NORMDIST(x;mean;standard\_dev;cumulative)$ , in our case being  $NORMDIST(t;m;s;FALSE)$ , and more specifically

$NORMDIST(\text{column B}; \text{cell K1}; \text{cell K3}; FALSE)$ .

- In column G, multiply the values in column F with the parameter  $T$ , i.e., column F\*cell K2, for days 1 to  $m$ .

- In column H, calculate the expected values  $E[Y]$  for days 1 to  $m$ , by multiplying the weekday effects  $w$  from column E to column G for each value/row, i.e., column G\*column E.

- In column I, calculate the probability  $p(Y = y)$  for days 1 to the last known day, under the assumption that  $Y_t \sim Poisson(T \times w \times f(t: m, s))$ , and using the expected values  $E[Y_t] = (T \times w \times f(t: m, s))$  from column H, by applying the Excel function  $POISSON.DIST(x;mean;cumulative)$ , in our case being  $POISSON.DIST(y; \text{column H}, FALSE)$ .

- In column J, the logarithmic values of column I are calculated (i.e.,  $\text{LOG}(\text{column I})$ ).

#### 4. Maximizing the likelihood:

- Start *Solver* in Excel. (Since Solver is not activated per default in Excel, it needs to be added first. To do so, click the *Microsoft Office Button* and then click *Excel Options*. Thereafter click *Add-ins* and then in the *Manage* box, select *Excel Add-ins* and click *Go*. In the *Add-Ins available* box, select the *Solver Add-in* check box and then click *OK*.)

- In *Set Objective*, choose cell K4 (being the value of interest) and choose to maximize it (*To: MAX*).

- In *By Changing Variable Cells*, choose cell K2 and K3. Note that appropriate starting values of K2 and K3 must be inserted manually first (see example of starting values above).

- Click the box in front of *Make Unconstrained Variables Non-Negative*.

- Run the solver by clicking on *Solve*.

- *Solver* will now estimate the values in K2 (the total number of ILI-cases of the whole epidemic) and K3 (The spread in time  $s$ ) in such way so that cell K4 (the sum of the logarithmic likelihood) is maximized.

5. The predicted peak intensity can now be read out on day  $m$  (the row of the peak timing day) in column H.