A systematic review estimated that the basic reproduction number ($R_0$) for coronavirus disease (COVID-19) is 2–3 (1). However, alone is insufficient to characterize an epidemic. The distribution of the serial interval (i.e., the length of time between symptom onset of 2 cases) has been estimated for COVID-19; mean intervals range from 3.1 to 7.5 days (2,3). Estimation of the generation interval ($T_g$) (i.e., the length of time between the points of infection for 2 linked cases) is less common. Although studies have reported means of 3.3 and 5.0 days (4; Li et al., unpub. data, https://doi.org/10.1101/2020.02.26.20028431), Ganyani et al. (5) estimated the mean ($\pm$ SD) of $T_g$ to be 3.9 ($\pm$ 2.7) days on the basis of which they estimated that 66% (95% credible interval [CrI] 45%–84%) of transmission occurred before symptoms. Another study of 77 pairs estimated the same proportion to be 44% (95% CI 25%–69%) (6). Because conventional outbreak control measures are centered around isolation, contact tracing, and treatment of symptomatic case-patients, a high prevalence of presymptomatic transmission ($p$) would warrant shifting measures to address potential transmission among persons with no apparent symptoms (7). Hence, to inform control measures for the outbreak in Singapore, we generated estimates of $T_g$, $R_0$, and $p$ by using published symptom onset data for COVID-19 cases in Singapore.

The Study
We implemented a cross-sectional study design to estimate $T_g$, $R_0$, and $p$ for the COVID-19 outbreak in Singapore during January 23–April 6, 2020. Given that containment measures were initiated over the duration of the study, we considered $R_0$ to be the effective reproduction number of the outbreak. All confirmed COVID-19 cases classified by the Ministry of Health of Singapore (MOH) as linked to a local cluster were included in this analysis. Information on case number, cluster, patient age and sex, imported status, date of symptom onset (DOO), and known contacts who have also been confirmed as case-patients were extracted from daily press releases published by MOH. DOOs for cases that were not available from press releases were extracted from a similar anonymized dataset of COVID-19 admissions to the National Centre for Infectious Diseases, Singapore. Cases with DOOs not available from that dataset were subsequently excluded from analysis. Our study was approved by the ethics review board of National Healthcare Group, Singapore.

We identified index cases and potential infectors of each case-patient on the basis of available information of the case-patients’ known contacts, published case links, and a heuristic to sensibly include potential infectors who could have transmitted the infection to the case-patients (Appendix, https://wwwnc.cdc.gov/EID/article/27/2/20-3018-App1.pdf). We subsequently used the infector–infectee pairs constructed to estimate the serial and generation interval distribution.
Assuming the same incubation period with mean (± SD) of 5.2 (± 2.8) days, we replicated the Bayesian Markov chain Monte Carlo procedure detailed in Ganyani et al. (5) to estimate the mean (SD) of the \( T_g \) (Appendix). With the estimated parameters, we constructed the distribution of \( R_0 \) and subsequently \( p \) by simulating infections and computing the proportion of presymptomatic transmissions. We conducted subgroup analyses for case-patients with multiple, family, or no contacts, and for clusters with no missing DOO. We conducted sensitivity analyses estimating the distribution of \( R_0 \) by using resampled values from a 95% CI of the epidemic growth rate and group-specific rates. For each distribution, we reported the median and 95% CrI. All analyses were conducted by using RStudio 1.2.5033 (https://rstudio.com).

A total of 1,375 confirmed cases had been reported as of April 6, 2020, and we applied our exclusion criteria to obtain a final sample size of 257 cases (Figure 1). We have summarized sample characteristics (Table 1) and the spread of cases over time (Figure 2). Because 48 index case-patients had no known infector, a maximum of 209 infector–infectee pairs were constructed for analysis.

Analyzing the 209 pairs, we estimated the mean \( T_g \) to be 3.44 (95% CrI 2.79–4.11) days, with an SD of 2.39 (95% CrI 1.27–3.45) days (Table 2). This estimate corresponded to an \( R_0 \) of 1.09 (95% CrI 1.08–1.11) and \( p \) of 0.72 (95% CrI 0.64–0.80). We estimated the serial interval distribution (Appendix Table 1) and convergence plots for all analyses (Appendix Figure 3).
Examining the 93 pairs with only 1 known contact, the estimates for mean $T_g$, SD $T_g$ and $R_0$ increased, whereas $p$ decreased (Table 2). The 116 pairs that required identification of potential infectors had a shorter mean $T_g$ and a higher $p$ in comparison (Table 2). Subgroup analyses are summarized in Appendix Table 2. However, the chains for pairs with family or no known contact exhibited poor convergence, and estimates were not reported. Sensitivity analyses using resampled growth rates and group-specific rates did not yield estimates differing from those of the main analyses (Appendix Table 3).

**Conclusions**

The mean generation interval of the COVID-19 outbreak in Singapore was estimated at 3.44 days, suggesting that an infected person would be expected to pass on an infection to another person in 3 days, within the range of 3.3–5.0 days reported by other studies (4,5; Li et al., unpub. data). Pairs with only 1 known contact yielded a larger estimate of 3.93 days, whereas pairs for whom infectors were identified had a shorter mean generation interval of 3.03 days. These results suggest that we might best report the upper bound of estimates, accounting for the presence of unclear transmission links within the clusters.

The $R_0$ estimated was slightly >1, higher than other estimates reported as of March 31, 2020 (8). We observed a high $p$, potentially a result of prompt isolation of symptomatic case-patients (M. Casey et al., unpub. data, https://doi.org/10.1101/2020.05.08.20094870). This higher proportion might also be attributable to our allowance of infector DOOs to be up to 3 days after their infectees’ DOOs, establishing the plausibility of presymptomatic transmission. We acknowledge that this cutoff would have an influence on our eventual estimates. Although negative serial intervals >3 days have occurred in other studies (5; Z. Du et al., unpub. data, https://doi.org/10.1101/2020.02.19.20025452), we chose a conservative cutoff of 3 days consistent with He et al. (6), where 9% of transmissions would occur before 3 days before DOO.

Nonetheless, the high prevalence of presymptomatic transmission in the community requires public health strategies to be responsive to this characteristic to remain effective. Universal wearing of masks in the community might reduce the likelihood of transmission through saliva and respiratory droplets (9). In place of testing when symptoms are observed,
universal testing of persons living in or working with confined populations should be prioritized to mitigate the risk for transmission of the infection into these populations (10). Contact tracing should be modified to include the period before symptom onset (6,7) and should adopt a digital approach to be more comprehensive and less labor intensive (4).

Our study generated estimates that accounted for the uncertainty arising from multiple potential infectors and a small sample size, which contributes to the scarce information about disease characteristics. Because we dropped cases without a reported DOO, and DOO data and contact information were self-reported, our estimates might be subject to selection, self-report, and recall biases. Our estimation approach assumed equal probability of infecting among potential infectors, although a higher likelihood of transmission among household contacts has been suggested (11). We also did not account for the potential formation of cyclical infector networks, although their effects on the estimates have been demonstrated to be limited (12). Nevertheless, our estimates contribute to knowledge about the transmission dynamics of COVID-19 and have implications for control measures.

Acknowledgments
We thank the Singapore Ministry of Health for their tireless efforts in outbreak control and publication of the data for this manuscript.

N.H. acknowledges funding provided by the EpiPose project from the European Union’s SC1-PHE-CORONA-VIRUS-2020 programme (project no. 101003688) and by the European’s Horizon 2020 research and innovation programme (grant agreement no. 682540–TransMID).

About the Author
Ms. Ng is a research analyst with the National Healthcare Group in Singapore. Her research interests include the application of biostatistics in epidemiology and health services research.

References
Appendix

**Identifying index cases within each cluster**

Within each cluster, an index case was either

(i) a primary case determined by epidemiologic investigations by the Ministry of Health, Singapore,

(ii) an imported case,

(iii) a case with the earlier date of symptom onset (DOO) in a cluster with only one other case,

(iv) a case with the earliest DOO in the cluster, and no subsequent cases with a DOO within 3 days after it.

The index cases would not have any possible infectors by definition. Clusters with no cases satisfying the criteria (i) to (iv) would not have a defined index case.

**Heuristic to identify potential infectors of cases**

We identified potential infectors of each case based on available information of the cases’ known contacts, published case links (1,2), and a heuristic to sensibly include a pool of candidates who could have transmitted the infection to the cases.

We defined a potential infector as an infector with a DOO within a period spanning 14 days before and 3 days after the DOO of the case. These thresholds were chosen to ensure that an infector-infectee pair would generate a plausible serial interval. We described two scenarios in Appendix Figure 1 to illustrate the implications of the thresholds. The first scenario would occur when the DOO of the infectee is after that of the infector (Appendix
Figure 1). As the maximum incubation period of COVID-19 has been suggested to be 14 days (3), we have considered the maximum plausible serial interval to be 14 days. The second scenario would occur when the DOO of the infectee is before that of the infector (Appendix Figure 1). As analyses by He et al found only 9% of transmission to occur 3 days before an infector’s DOO and recommended for contact tracing to include this window (4), the minimum serial interval we would consider would be −3 days.

We then constructed infector-infectee pairs of cases by assigning an infector to each case. Within clusters that have only two cases, the index case was assigned as the infector for the other case in the cluster. Within larger clusters, cases that have only one known contact were assigned this contact as their infector (Appendix Figure 2).

For cases with multiple contacts, a potential infector was randomly selected from the pool of contacts using an independence sampler in the Markov Chain Monte Carlo (MCMC) algorithm. For cases with no known contacts, a potential infector was similarly selected from other cases within the same cluster. If a case had no potential infectors with a DOO within the period for plausible serial intervals, the case’s known contacts were nonetheless assigned as potential infectors.

**Estimating the distributions of transmission parameters**

Assuming the same incubation period with mean of 5.2 days and standard deviation of 2.8 days (5), we replicated the Bayesian MCMC procedure as detailed in Ganyani et al (6), to estimate the mean and standard deviation of the generation interval \( T_g \) distribution. Briefly, the infectors assigned and parameters of the \( T_g \) distribution were updated in a two-step process. The unknown infector-infectee links were updated using an independence sampler. Uniform priors were assigned to the parameters of the generation interval distribution and updated using a random-walk Metropolis-Hastings algorithm with a uniform proposal distribution (7). The posterior distribution was modeled using 3,000,000 iterations of which the first 500,000 were discarded as burn-in, thinning applied every 200 iterations, resulting in 12,500 iterations of accepted parameter estimates.

With the 12,500 instances of estimates and the spread of cases, we computed the corresponding effective \( R_0 \) estimates using the relationship \( R_0 = e^{r\mu - \frac{1}{2}r^2\sigma^2} \), where \( r \) is the growth rate of the epidemic and \( \mu, \sigma \) are the mean and standard deviation of the generation interval distribution (8). Using the same parameter estimates and the incubation period
parameters assumed above, we simulated 1000 transmissions per iteration and computed the pre-symptomatic proportion as the proportion of transmissions that occurred before the end of the infector’s incubation period. For each distribution, the median and 95% credible interval were reported.

References


### Appendix Table 1. Estimated serial interval (SI) distributions*

<table>
<thead>
<tr>
<th>Type of infector–infectee pairs</th>
<th>Median (95% credible interval)</th>
</tr>
</thead>
<tbody>
<tr>
<td>All cases (N = 209)</td>
<td>Simulated SI (−5.41, 13.11)</td>
</tr>
<tr>
<td>Cases with only 1 known contact (N = 93)</td>
<td>3.74 (−5.10, 13.91)</td>
</tr>
<tr>
<td>Cases with multiple or no contacts (N = 116)</td>
<td>2.83 (−5.83, 12.83)</td>
</tr>
</tbody>
</table>

*SI, Serial interval; DOO, Date of symptom onset.

### Appendix Table 2. Subgroup analyses*

<table>
<thead>
<tr>
<th>Infectee type</th>
<th>Mean $T_g$</th>
<th>SD $T_g$</th>
<th>$R_0$</th>
<th>$p$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Cases with multiple known contacts (N = 72)</td>
<td>3.72 (2.57, 5.04)</td>
<td>3.22 (1.49, 5.43)</td>
<td>1.10 (1.07, 1.14)</td>
<td>0.69 (0.56, 0.81)</td>
</tr>
<tr>
<td>Cases in clusters with no missing DOOs (N = 103)</td>
<td>4.05 (3.13, 5.06)</td>
<td>2.90 (1.36, 4.42)</td>
<td>0.98 (0.97, 0.98)</td>
<td>0.64 (0.53, 0.74)</td>
</tr>
</tbody>
</table>

$T_g$, generation time; SD, standard deviation; $R_0$, basic production number; $p$: pre-symptomatic proportion; DOO: Date of symptom onset.

### Appendix Table 3. Sensitivity analyses of reproduction number $R_0$ distribution

<table>
<thead>
<tr>
<th>Infectee type</th>
<th>No. time points</th>
<th>Max daily incident cases</th>
<th>Growth rate of log (cases)</th>
<th>Median (95% credible interval)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Main analyses</td>
<td></td>
<td></td>
<td></td>
<td>Original $R_0$, Revised $R_0$</td>
</tr>
<tr>
<td>All cases (N = 209)</td>
<td>37</td>
<td>13</td>
<td>0.027</td>
<td>1.09 (1.08, 1.11), 1.90 (1.03 to 1.17)*</td>
</tr>
<tr>
<td>Cases with only 1 known contact (N = 93)</td>
<td>32</td>
<td>7</td>
<td>0.017</td>
<td>1.11 (1.08, 1.14), 1.07 (1.05, 1.09)</td>
</tr>
<tr>
<td>Cases with multiple or no known contact (N = 116)</td>
<td>35</td>
<td>13</td>
<td>0.021</td>
<td>1.08 (1.06, 1.11), 1.06 (1.04, 1.08)</td>
</tr>
<tr>
<td>Subgroup analyses</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cases with multiple known contacts (N = 72)</td>
<td>8</td>
<td>5</td>
<td>0.104</td>
<td>1.10 (1.07, 1.14), 1.39 (1.22, 1.57)</td>
</tr>
<tr>
<td>Cases in clusters with no missing DOOs (N = 103)</td>
<td>34</td>
<td>7</td>
<td>−0.003</td>
<td>0.98 (0.97, 0.98), 0.99 (0.99, 0.99)</td>
</tr>
</tbody>
</table>

*Estimates computed using resampled growth rates.

### Appendix Figure 1. Time intervals within a transmission chain.
Appendix Figure 2. Heuristic for identifying potential infectors of a case.

Appendix Figure 3. Convergence plots.